# Firm Level Pass Through: A Machine Learning Approach

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#### Introduction

- A consistent finding of the exchange rate pass through (ERPT) literature is that export prices denominated in destination currency do not react one-to-one to bilateral exchange rate movements.
- Among many theoretical and empirical studies, one approach is to use micro-level firm data to understand ERPT.
- This paper proposes an innovative method to systematically detect features affecting ERPT and predict ERPT at the firm-level.
- The proposed method works with micro-founded macro models and is directly applicable to highly dis-aggregated large scale custom datasets.

## Exchange Rate Pass Through

The price of an exporter denominated in the destination currency at the border can be expressed as

$$p_{dt}^{destination} = MKP_{dt} \cdot e_{dt} \cdot MC_t$$

- *MKP*<sub>dt</sub> is the exporter's markup at destination d.
- $e_{dt}$  is the bilateral exchange rate between the origin country and destination d ( $\uparrow$  means origin appreciation).
- *MC<sub>t</sub>* is the marginal cost of the firm.

Alternatively, write into the price denominated in the origin currency.

$$p_{dt}^{origin} = MKP_{dt} \cdot MC_t$$

Exchange rate pass through to the import price at the destination d is defined as

$$eta_{dt} := rac{\partial log(p_{dt}^{destination})}{\partial log(e_{dt})}$$

# **Empirical Findings**

#### Macro level (the puzzle):

- **1** ERPT is far from complete;
  - Campa and Goldberg (2005) estimate  $\beta = 0.46$  [one month] and  $\beta = 0.64$  [4 months] for OECD countries.
- Ø different across industries and countries;
- 3 changes over different time horizons;
  - Marazzi et al. (2005); Gust, Leduc and Vigfusson (2010)

Micro firm-level studies addressing the puzzle find evidences cannot be easily reconciled with macro findings:

- 1 more complete ERPT using unit values and firm level dataset.
  - $\beta = 1 0.08 = 0.92$ , Berman, Martin, and Mayer (2012), France.
  - $\beta = 1 0.20 = 0.80$ , Amiti, Itskhoki, and Konings (2014), Belgium.
  - $\beta = 1 0.05 = 0.95$ , Corsetti, Crowley, Han, and Song (WP), China.
- 2 Lewis (2016) discusses the reconcilability of ERPT estimates for UK import prices.

# Missing Components

#### 1 The source of the exchange rate shock matters.

- Theory: Corsetti, Dedola, and Leduc (2008)
- Empirical: Forbes, Hjortsoe, Nenova (2015)

#### Ø Heterogeneity in ERPT

- Competition and market structure matters. ERPT is U-shaped in market share: Dornbusch (1987), Atkeson and Burstein (2008), Auer and Schoenle (2016)
- Other stuff matters. Lower ERPT if more imported inputs; higher productivity; higher distribution margin; low quality; invoiced in international currencies; differentiated products; general trade v.s. processing materials; trading v.s. non-trading exporters;
- 3 Not only bilateral but also multi-lateral exchange rates.
  - Trade flows: Bown and Crowley (2007)
- 4 Nominal rigidities.
  - Conditional on price changes: Gopinath and Itskhoki (2010), Fitzgerald and Haller (2013)

### A method customized for big data

- Unlike micro studies in other fields, international trade firm-level datasets recently made available contain a significant portion of firms in an economy and almost all custom transactions at firm product (8-digits) level in a given period.
- Although significant progress has been made in understanding the heterogeneity of ERPT, the literature fails to make scenario dependent predictions based on the macro environment and observed firm, product and market structure features.

This paper tries to construct a method to systematically detect features affecting ERPT and predict ERPT at the firm-level.

# Challenges

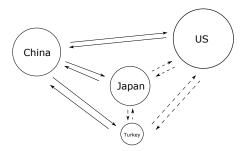


Figure 1 A multi-country trade model with multilateral exchange rate shocks

- 1 Unobserved variables
  - marginal cost of firms
  - trade flows between trade partners of China
- 2 Nonlinearity
  - ERPT may be a nonlinear function of firm characteristics and market structures, e.g. U-shaped in market share.

# This paper

- Proposes an innovative method that estimates determinants of ERPT at the firm-level accounting for
  - nonlinearity
  - unobserved variables not varying at all dimensions.
- Extends the two-country Atkeson and Burstein (2008) model into a multi-country framework and studies how markets reach equilibrium under multilateral exchange rate shocks.
- Using the simulated data from my model, I test the accuracy of the method in identifying ERPT at the firm-level.
- Applying my method with China's custom data, I document evidences on the nonlinear relationship between market shares and ERPT.

### Outline

#### 1 Question

- 2 Algorithm
- 3 Tested on a Multi-country Atkeson and Burstein (2008) Model
- 4 Conclusions

#### A general pricing equation for exports

$$p_{i,f,d,t} = g(\mathcal{F}_{i}, \mathcal{F}_{f}, \mathcal{F}_{d}, \mathcal{F}_{t}, \mathcal{F}_{f,d}, \mathcal{F}_{f,t}, \mathcal{F}_{f,t}, \mathcal{F}_{d,t}, \mathcal{F}_{f,d,t}, \mathcal{F}_{f,d,t}, \mathcal{F}_{i,f,t}, \mathcal{F}_{i,f,d}, \mathcal{F}_{i,f,d,t}, \mathcal{F}_{i,f,d,t}, \mathcal{F}_{i,f,d,t}, \mathcal{F}_{i,f,d,t})$$

Where i for product, f for firm, d for destination.

- $\mathcal{F}_i$ : product specific factors/properties, etc.
- $\mathcal{F}_{d,t}$ : exchange rate, CPI, and other macro variables.
- $\mathcal{F}_{i,f,t}$ : marginal cost, etc.
- $\mathcal{F}_{f,d,t}$ : firm destination market share, etc.
- $\mathcal{F}_{i,f,d,t}$ : firm-product destination market share, etc.

### Question

$$p_{\mathcal{I}} = g\left(\underline{e_{d,t}}, X_{\mathcal{I}}, \underline{M_{\mathcal{Q}}}, \epsilon_{\mathcal{I}}\right)$$

where

- $\mathcal{I} = \{i, f, d, t\}; \mathcal{Q} \subset \mathcal{I}$
- g: an unknown function
- *p*: the exporter's price
- e: the bilateral exchange rate
- X: a vector of observed feature variables
- M: a vector of unobserved variables that correlate with e
- $\epsilon$ : the error term that does not correlate with e

I

Objective:

dentify 
$$\frac{\partial g(.)}{\partial e_{d,t}}$$

### Outline

# Question Algorithm Tested on a Multi-country Atkeson and Burstein (2008) Model Conclusions

#### An algorithm to learn the behavior of exporters

- Machine Learning in Economics:
  - Hal R. Varian introduced various machine learning methods in 2014.
  - Athey and Imbens (2015), Bajari, et al. (2015), Chernozhukov, Hansen, and Spindler (2015), Kleinberg, et al. (2015), Wager and Athey (2015), Athey and Imbens (2016), Chernozhukov, et al. (2016)
- Gradient Boosting Regression Trees (GBRT):
  - Recent applications of GBRT:
    - Bhatt, et al. (2013, Nature); Pearson, et al. (2014, Nature Climate Change); Doench, et al (2016, Nature Biotechnology), etc.
  - High precision in solving practical estimation problems:
    - the main estimation method of winners of many international data science competitions, e.g. Kaggle competitions.
- Critiques of machine learning methods:
  - Completely data driven
  - Good at making predictions but does not identify causal relationships and enhance our economic understanding.

#### My contribution to the ML literature

- The proposed algorithm represents an attempt to control unobserved variables by feeding additional structural information implied by economic models into the machine learning algorithm.
  - no longer completely data driven but relies on the supplied external structural information.
- The core:
  - I show a unique property of tree based algorithms that can be exploited to control unobserved components.
  - In a multi-dimensional panel, parameter estimates from a structural model in a range of limited-dimensional spaces can help to restrain the variation of unobserved components.

1 For  $t = 1...n_t$ , run OLS, and collect coefficients  $b_t^0$ ,  $b_t^1$ 

$$p_{d,t} = b_t^0 + b_t^1 e_{d,t}$$

2 For  $d = 1...n_d$ , run OLS, and collect coefficients  $b_d^0$ ,  $b_d^1$ 

$$p_{d,t} = b_d^0 + b_d^1 e_{d,t}$$

3 Approximating p. Run gradient boosting regression tree algorithm entering coefficients {b<sub>t</sub><sup>0</sup>, b<sub>t</sub><sup>1</sup>, b<sub>d</sub><sup>0</sup>, b<sub>d</sub><sup>1</sup>} as additional feature variables. Obtain

$$g_1: (e_{d,t}, m_{s_{d,t}}, b_t^0, b_t^1, b_d^0, b_d^1) \to p_{d,t}$$

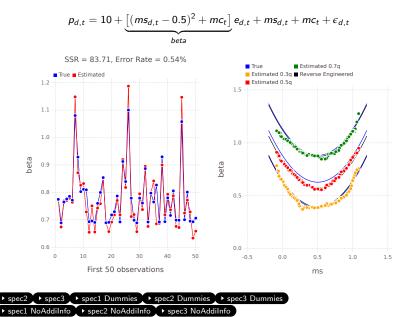
Wumerical differentiation. Use g<sub>1</sub> to construct counter-factual predictions conditional on the values of ms<sub>d,t</sub>, b<sup>1</sup><sub>t</sub>, b<sup>1</sup><sub>t</sub>, b<sup>0</sup><sub>d</sub>, b<sup>1</sup><sub>d</sub> and calculate:

$$\begin{array}{l} \bullet \ p_{d,t}^{\text{Est1}} = g_1(e_{d,t}, m_{sd,t}, b_t^0, b_t^1, b_d^0, b_d^1) \\ \bullet \ p_{d,t}^{\text{Est2}} = g_1(e_{d,t} + std(e_{d,t}), m_{sd,t}, b_t^0, b_t^1, b_d^0, b_d^1) \\ \bullet \ \beta_{d,t}^{\text{Est2}} = \frac{p_{d,t}^{\text{Est2}} - p_{d,t}^{\text{Est1}}}{std(e_{d,t})} \end{array}$$

S Approximating β<sup>Est</sup>. Run GBRT with the dependent variable β<sup>Est</sup><sub>d,t</sub> on e<sub>d,t</sub>, ms<sub>d,t</sub>, b<sup>0</sup><sub>t</sub>, b<sup>1</sup><sub>t</sub>, b<sup>0</sup><sub>d</sub>, b<sup>1</sup><sub>d</sub>, and get

$$g_2: (e_{d,t}, \textit{ms}_{d,t}, b^0_t, b^1_t, b^0_d, b^1_d) 
ightarrow eta^{\textit{Est}}_{d,t}$$

#### The Proposed Algorithm



#### Outline

# Question Algorithm Tested on a Multi-country Atkeson and Burstein (2008) Model Conclusions

# A Multi-country Atkeson and Burstein (2008) Model

- Three countries
- Market structure assumptions as in Atkeson and Burstein (2008)
- Large number of sectors and N firms competing in each sector
- Only the best firm in each sector exports.
- N-2 number of destination competitors
- Productivity distributions are different across sectors, countries.
- Simplification:
  - No imported inputs
  - No nominal rigidities
  - No distribution cost
- Details of Model Specifications

#### Recover ERPT for Simulated Exporters

- Simulate the model for 240 periods (20 years).
- Construct an environment similar to my empirical dataset:
  - 1 Select an origin country
  - 2 Record trade flows of exports

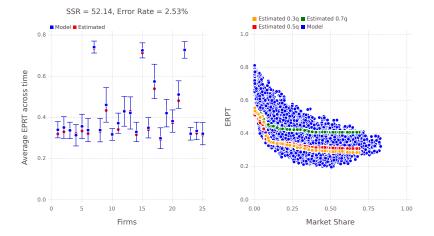
 $f, s, d, t, p_{f,s,d,t}, q_{f,s,d,t}, P_{s,d,t}, D_{s,d,t}, e_{d,t}, P_{d,t}, L_{d,t}, C_{d,t}$ 

Objective: Use only information from recorded variables to

- Learn trade and pricing patterns of exporters.
- Q Given market conditions at t 1, estimate price changes under a bilateral exchange rate shock at period t.
- 8 Recover ERPT for simulated firms.

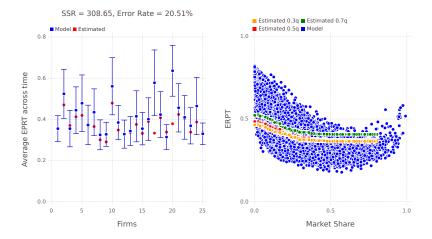
## Case 1: Only Exchange Rate Shocks

My algorithm compared to true counter-factual environments



#### Case 2: Add productivity shocks

My algorithm compared to true counter-factual environments



### Outline

# Question Algorithm Tested on a Multi-country Atkeson and Burstein (2008) Model Conclusions

#### Conclusions

- Proposes an innovative method that works with micro-founded macro models and is directly applicable to highly dis-aggregated large scale custom datasets. It estimates determinants of ERPT at the firm-level accounting for
  - nonlinearity
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# Appendix

#### Outline

#### **5** Appendix

Model Classification and Regression Trees One Dimensional Examples Drawbacks of Conventional Regression Approaches Alternative Specifications Full Algorithm

#### Firm's Problem

Cournot competition in quantities within sector *s*:

$$\max_{q_{f,s,o,d,t}} q_{f,s,o,d,t} (p_{f,s,o,d,t} - mc_{f,s,o,t})$$
(1)

subject to

$$q_{f,s,o,d,t} = \left(\frac{p_{f,s,o,d,t}}{P_{s,d,t}}\right)^{-\rho_s} \left(\frac{P_{s,d,t}}{P_{d,t}}\right)^{-\eta} D_{d,t}$$
(2)

$$D_{s,d,t} \equiv \left[\sum_{o} \sum_{f \in s, f \in I_{o}} (q_{f,s,o,d,t})^{\frac{\rho_{s-1}}{\rho_{s}}}\right]^{\frac{\rho_{s}}{\rho_{s-1}}} , \quad D_{d,t} \equiv \left[\sum_{s} (D_{s,d,t})^{\frac{\eta-1}{\eta}}\right]^{\frac{\eta}{\eta-1}}$$
$$P_{s,d,t} \equiv \left[\sum_{o} \sum_{f \in s, f \in I_{o}} (p_{f,s,o,d,t})^{1-\rho_{s}}\right]^{\frac{1}{1-\rho_{s}}} , \quad P_{d,t} \equiv \left[\sum_{s} (p_{s,d,t})^{1-\eta}\right]^{\frac{1}{1-\eta}}$$

where f, s, o, d, t = firm, sector, origin, destination, time

#### Price, Market Share and Demand Elasticity

The usual pricing equation

$$p_{f,s,o,d,t} = \frac{\varepsilon_{f,s,o,d,t}(ms_{f,s,o,d,t})}{\varepsilon_{f,s,o,d,t}(ms_{f,s,o,d,t}) - 1} \frac{mc_{f,s,o,t}}{e_{d,t}}$$

Demand elasticity  $\varepsilon$  is a nonlinear function of market share.

$$\varepsilon = \frac{1}{\frac{1}{\rho}(1 - ms) + \frac{1}{\eta}ms}$$

where

$$ms = rac{pq}{\sum_f pq} = rac{p^{1-
ho}}{\sum_f (p)^{1-
ho}}$$

Subscripts f, s, o, d, t is omitted for simplicity.

#### Equilibrium Effect of Market Share Matters

$$\widehat{p}_{k,s,o,d,t} = \kappa_{k,s,o,d,t} \widehat{ms}_{k,s,o,d,t} + \widehat{mc}_{k,s,o,t} + \widehat{e}_{o,d,t}$$

where  $\kappa_{f,s,o,d,t}$  is the price elasticity with respect to a firm's own market share:

$$\kappa_{f,s,o,d,t} \equiv \left(\frac{\varepsilon_{f,s,o,d,t}}{\varepsilon_{f,s,o,d,t}-1}\right) \left(-\frac{1}{\rho_s} + \frac{1}{\eta}\right)$$

Equilibrium effects of market share under a multi-country context:

$$\begin{split} & \frac{\widehat{ms}_{k,s,o,d,t}}{\lambda_{f,s,o,d,t}} = (1 - ms_{k,s,o,d,t}) \left\{ (1 - \rho_s) \left[ \widehat{mc}_{k,s,o,t} + \widehat{e}_{o,d,t} \right] \right\} \\ & - \sum_{o'} \sum_{f \neq k} ms_{f,s,o',d,t} \left\{ (1 - \rho_s) \left[ \widehat{mc}_{f,s,o',t} + \widehat{e}_{o',d,t} - \kappa_{f,s,o',d,t} \widehat{ms}_{f,s,o',d,t} \right] \right\} \end{split}$$

where  $\lambda_{k,s,o,d,t}$  is the theoretical exchange rate pass through and it is U-shaped in market share:

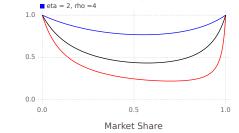
$$\lambda_{f,s,o,d,t} = \frac{1}{1 - (1 - ms_{f,s,o,d,t})(1 - \rho_s)\kappa_{f,s,o,d,t}}$$

#### The Theoretical U-shape and More

$$\widehat{p}_{k,s,o,d,t} = \lambda_{k,s,o,d,t} \left[ \widehat{mc}_{k,s,o,t} + \widehat{e}_{o,d,t} - \kappa_{k,s,o,d,t} \widehat{CE}_{k,s,o,d,t} \right]$$

where  $\widehat{CE}_{k,s,o,d,t}$  is the total effect of competitors' reactions.

$$\widehat{CE}_{k,s,o,d,t} = \sum_{o'} \sum_{f \neq k} s_{f,s,o',d,t} (1 - \rho_s) \left[ \widehat{mc}_{f,s,o',t} + \widehat{e}_{o',d,t} - \kappa_{f,s,o',d,t} \widehat{ms}_{f,s,o',d,t} \right]$$



eta = 2, rho =10 eta = 1.2, rho =10

The theoretical U-shape of  $\lambda \rightarrow$ 

$$\lambda = \frac{1}{1 - \frac{(1-ms)(1-\rho)ms}{1 - \frac{1-ms}{\rho} - \frac{ms}{\eta}}(1/\eta - 1/\rho)}$$

#### Household's Problem

$$\max_{C_{d,t},L_{d,t}} E_0 \sum_{t=0}^{\infty} \beta^t U(C_{d,t}, L_{d,t})$$

subject to

$$U_{d,t} = \log[C_{d,t}^{\mu}(1 - L_{d,t})^{1-\mu}]$$

As in Atkeson and Burstein (2008), household can trade a complete set of international assets.

$$P_{d,t}C_{d,t} + \sum_{o} \left[ \sum_{v} p_{o,t}^{B}(v) B_{o,t}(v) - (1 + i_{o,t-1}) B_{o,t-1} \right] * e_{o,d,t} = W_{d,t}L_{d,t} + \Pi_{d,t}$$

The optimal solution of household's problem are given by

$$\frac{1-\mu}{\mu}\frac{C_{d,t}}{1-L_{d,t}} = \frac{W_{d,t}}{P_{d,t}}$$
(3)

$$\frac{C_{o,t}P_{o,t}}{e_{o,d,t}C_{d,t}P_{d,t}} = \frac{C_{o,t+1}(\nu)P_{o,t+1}(\nu)}{e_{o,d,t+1}(\nu)C_{d,t+1}(\nu)P_{d,t+1}(\nu)}$$
(4)

#### Other Equilibrium Conditions

$$mc_{f,s,o,t} = \frac{W_{o,t}}{\Omega_{f,s,o,t}}$$
$$\Omega_{f,s,o,t}l_{f,s,o,t} = \sum_{d} q_{f,s,o,d,t}$$
$$\sum_{f,s} l_{f,s,o,t} = L_{o,t}$$

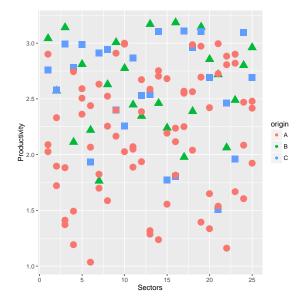
Steady state condition: balance of trade

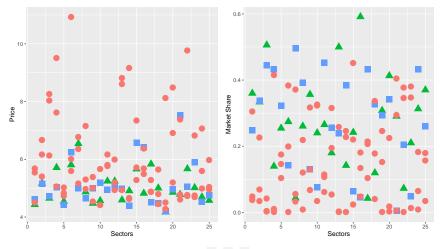
$$\sum_{f,s} p_{f,s,d,o} q_{f,s,d,o} = \sum_{f,s} p_{f,s,o,d} q_{f,s,o,d} * e_{o,d} \quad \text{for } o \neq d$$

where  $e_{o,d}$  is defined as units of currency o per units of currency d.

	Countries	5	Ν	ρ	η	$\Phi(\Omega)$
Benchmark	3	25	3+2	10	2	Uniform

#### Distribution of Simulated Productivities in Country A

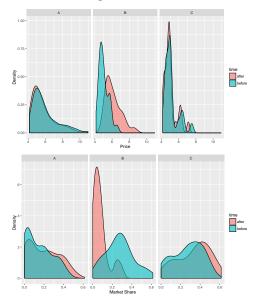




origin 🔹 A 🔹 B 🔹 C

# Appreciation of Country B

Change in Price and Market Share Dist in A



#### **Estimation Procedure**

 Make assumptions based on the structural model. In general, the structural estimation can take any form. Here, I will use the linear approximation of the pricing equation:

$$log(p_{f,s,d,t}) = a + b * log(e_{d,t}) + c * log(p_{f,s,d,t-1})$$

- 2 Identify dimensions to be fixed as {s, t, d, sd}. Run regressions and collect coefficients a<sub>s</sub>, b<sub>s</sub>, c<sub>s</sub>, a<sub>t</sub>, b<sub>t</sub>, c<sub>t</sub>, ...
- **3** Create market share measure  $ms_{fsdt,sdt} = \frac{p_{f,s,d,t}q_{f,s,d,t}}{P_{s,d,t}D_{s,d,t}}$
- ④ Run GBRT with the dependent variable log(p<sub>f,s,d,t</sub>) on log(e<sub>d,t</sub>) and feature variables *F*

$$\begin{split} \mathcal{F} &:= & \log(e_{d,t-1}), \log(e_{d',t}), \log(e_{d',t-1}), \log(ms_{fsdt-1,sdt-1}), \\ & \log(D_{s,d,t-1}), \log(P_{s,d,t-1}), \log(P_{d,t-1}), \log(L_{d,t-1}), \log(C_{d,t-1}), \\ & a_s, b_s, c_s, a_t, b_t, c_t, \ldots \end{split}$$

and obtain g1

### Estimation Procedure (cont.)

S Numerical differentiation on predicted counter-factual bilateral exchange rates

$$\begin{split} p_{f,s,d,t}^{Est1} &= g_1[log(e_{d,t}), \mathcal{F}] \\ p_{f,s,d,t}^{Est2} &= g_1[log(e_{d,t-1}), \mathcal{F}] \\ ERPT_{f,s,d,t}^{Est} &= \frac{log(p_{f,s,d,t}^{Est1}) - log(p_{f,s,d,t}^{Est2})}{log(e_{d,t}) - log(e_{d,t-1})} \end{split}$$

**6** Run GBRT again with supervisor  $ERPT_{f,s,d,t}^{Est}$  on  $log(e_{d,t})$ ,  $\mathcal{F}$  and obtain  $g_2$ .

# Case 1: Only Exchange Rate Shocks

 No financial market exchange rate arbitrage condition holds

$$e_{1,2,t} = \frac{e_{1,3,t}}{e_{2,3,t}}$$

 Max 2 shocks (the third bilateral exchange rate is determined by the no arbitrary condition)

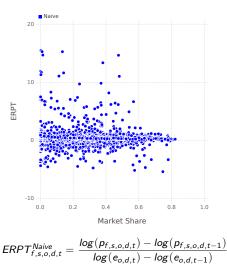
$$e_{1,2,t} = \xi_{1,t} e_{1,2,ss}$$
  

$$e_{3,2,t} = \xi_{3,t} e_{3,2,ss}$$
  

$$\xi_{i,t} \sim U(0.8, 1.2)$$

 Counter-factual state if there was no exchange rate shocks between 1 and 2

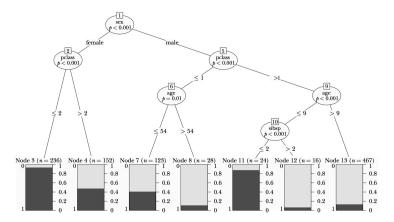
$$\begin{split} e^{c}_{1,2,t} &= e_{1,2,t-1} \\ e^{c}_{3,2,t} &= \xi_{3,t} e_{3,2,ss} \\ e^{c}_{1,3,t} &= \frac{e^{c}_{1,2,t}}{e^{c}_{3,2,t}} \end{split}$$



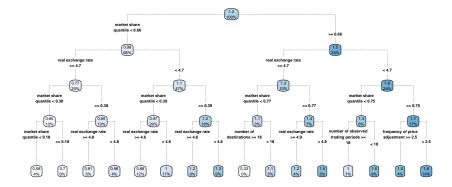
### Illustration: Classification and Regression Trees

A ctree for Survivors of the Titanic

(black bars indicate fraction of the group that survived)

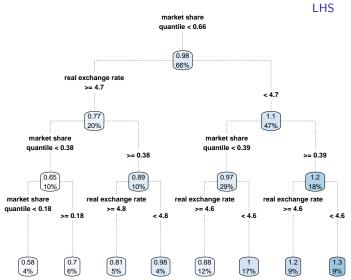


### Classification and Regression Trees Predicting Export Price



Source: my own calculation from China's import and export database.

# Classification and Regression Trees



#### **Classification and Regression Trees** RHS >= 0.66 1.5 34% real exchange rate >= 4.7 < 4.7 1.3 1.6 10% 24% market share market share quantile < 0.77 quantile < 0.75 >= 0.77 >= 0.75 1.4 1.1 3% 7% 6% .17% number of real exchange rate number of observed frequency of price destinations >= 18 trading periods >= adjustment >= 2.5 >= 4.9 < 18 < 4.9 < 2.5 18 < 18 0.33 1.1 1.5 1.5 1.2 1.5 0% 3% 4% 3% 1% 6% 4% 14%

# Gradient Boosting Models

Based on Friedman (2001), to minimize the objective function

$$\widehat{f}(\mathbf{x}) = \arg\min_{f(\mathbf{x})} E_{y,\mathbf{x}} \Psi(y, f(\mathbf{x}))$$
(5)

- **1** a loss function (distribution  $\Phi$ )
- 2 the number of iterations, iter
- 3 the depth of each tree, inter.depth
- 4 the shrinkage or the learning rate, lr
- **5** sampling rate (bagging fraction), bf

# Gradient Boosting Models

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	Φ	iter	inter.depth	lr	bf
Benchmark	Normal	Cross Validation	8	0.005	0.5
Robustness	Laplace, Quantile	5000	1-10	0.01, 0.001	0.3, 1

# Gradient Boosting Models (cont.)

Initialize  $\hat{f}(\mathbf{x})$  to be a constant. For *i* in 1,..., *Iter<sub>max</sub>* 

Compute the negative gradient as the working response

$$h_{i} = -\frac{\partial}{\partial f(\mathbf{x}_{i})} \Psi(y_{i}, f(\mathbf{x}_{i})) \Big|_{f(\mathbf{x}_{i}) = \widehat{f}(\mathbf{x}_{i})}$$
(6)

- 2 Randomly select a fraction bf from the dataset (Random Forest/Bagging)
- 3 Fit a regression tree with inter.depth splits, g(x), predicting h<sub>i</sub> from the covariates x<sub>i</sub>.
- 4 Update the estimate of  $f(\mathbf{x})$  as

$$\widehat{f}(\mathbf{x}) \to \widehat{f}(\mathbf{x}) + \ln * g(\mathbf{x})$$
 (7)

- 6 Repeat step 1-4 until Itermax
- 6 Cross validation method to determine the optimal iter

#### Back

# One dimensional example

Consider the case of identifying the individual treatment effect.

$$y_i = eta_i(M_i)D_i + M_i$$
  
 $eta_i(M_i) := M_i$   
 $D_i \in \{0, 1\}, \ M_i \in \{0, 1\}$ 

- where D<sub>i</sub> is a treatment indicator and β<sub>i</sub> is the treatment effect for individual *i*.
- The objective is to find β<sub>i</sub> given data of individual outcomes y<sub>i</sub> and its treatment indicator D<sub>i</sub>.
- The data generating process (the functional form of each variable) is unknown. *M<sub>i</sub>* is unobserved.

Suppose  $M_i$  is observed, we can estimate the individual treatment effect  $\beta_{it}$  using the following two-step procedure:

**1** Use a nonparametric econometric method or a machine learning algorithm to recognize the pattern of  $y_i$  using  $D_i$  and  $M_i$ . Obtain

$$g_1:(D_i,M_i)\to y_i$$

2 Use  $g_1$  to construct counter-factual predictions conditional on the value of  $M_i$  and calculate individual treatment effect.

$$\beta_i^{Est} = g_1(1, M_i) - g_1(0, M_i)$$

- In most cases, we do not observe  $M_i$ . But it may be possible to have/create a variable  $\mathfrak{M}_i$  that preserves some structural information of  $M_i$ .
- If we could construct counter-factuals conditional on the structural information provided by  $\mathfrak{M}_i$ , we will be able to recover  $\beta_i$  using the above procedure.
- In general, the structural information contained by the alternative variable  $\mathfrak{M}_i$  could be highly nonlinear.
- I find decision tree based algorithms have a unique advantage in addressing this type of problems.

# One dimensional example (cont.) Simulate 200 individuals:

$$y_i = \beta_i(M_i)D_i + M_i$$
  
$$\beta_i(M_i) = M_i$$

Table 2: Values of  $y_i$ 

1. 7	ssignment of <i>M</i> ;	Уi	βi	Mi	Di
Mi	i	0	0	0	0
)	1-100	0	0	0	1
	101-200	1	1	1	0
		2	1	1	1

• the assignment of  $M_i$  is constructed to be ordered

ŀ

 I want to utilize the information provided by the index *i* to estimate the individual treatment effect β<sub>i</sub>.

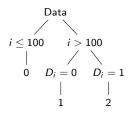
# A Basic Classification and Regression Tree (CART) Algorithm

A CART algorithm recursively binary splits/partitions data at the point which minimizes the mean prediction error (MPE) measured by the loss function h(.)

$$MPE = \sum_{\tau \in leaves(T)} \sum_{i \in \tau} h(y - m_c)$$
$$m_c = \frac{1}{n_c} \sum_{i \in \tau} y_i$$

- **1** The algorithm starts a tree of single node containing all points. If all the points in the node have the same value for all the input variables, stop.
- 2 Search over all binary splits of all variables for the one which will reduce MPE as much as possible. If the largest decrease in MPE would be less than some threshold, or one of the resulting nodes would contain less than q points, stop. Otherwise, take that split, creating two new nodes.
- **3** In each new node, back to step 1.

## Results



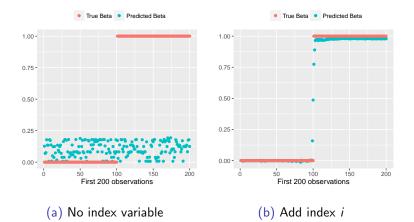
\_

i	$\widehat{y_i}$ Evalution $D_i = 1$	nated at $D_i = 0$	Estimated $\beta_i$	True $\beta_i$
1-50	0	0	0	0
51-100	Ō	0	0	Ō
101-150	2	1	1	1
151-200	2	1	1	1

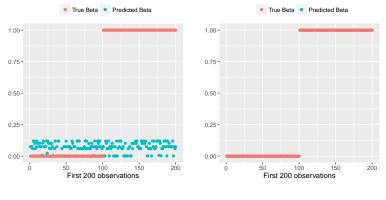
Stage 1

Stage 2

## Example 1: the ordered case



# Example 1: the ordered case (cont.)



(c) Add dummy variables of *i* 

(d) Add true  $\beta_i$ 

# The weak monotonic transformation property

### Definition 1

Let  $\{a_n\}$  be a sequence of real numbers.  $\exists: \{a_n\} \to \{b_n\} \in \mathbb{R}^N$  is a weak monotonic transformation if

$$egin{aligned} \mathsf{a}_{j} > \mathsf{a}_{i} \Rightarrow \mathsf{b}_{j} > \mathsf{b}_{i} & orall i, j \in \{1, ..., \mathsf{N}\} & ext{or} \ \mathsf{a}_{j} > \mathsf{a}_{i} \Rightarrow \mathsf{b}_{j} < \mathsf{b}_{i} & orall i, j \in \{1, ..., \mathsf{N}\} \end{aligned}$$

### **Proposition 1**

For a large number of observations n, entering  $\{a_n\}$  as a feature variable is equivalent to entering any  $\neg (\{a_n\})$  in decision tree based algorithms.

### Proposition 2

Let  $X_n$  be a set of feature variables excluding  $a_n$ . If  $var(a|X_n) \neq 0$  for some values of  $X_n$  and there is a large number of observations for these subsets of  $X_n$ , entering  $\{a_n\}$  as a feature variable is equivalent to entering any  $\exists (\{a_n\}|X_n)$  in a gradient boosting regression algorithm.

# Example 2: Weak and conditional weak monotonic transformation

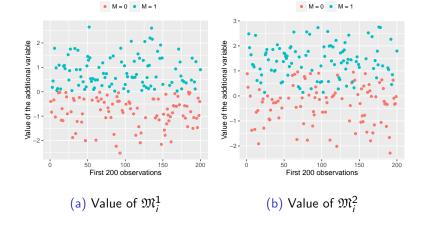
$$y_i = \beta_i(X_i, M_i)D_i + M_i$$
  
 $\beta_i(X_i, M_i) = X_i + M_i$   
 $D_i \in \{0, 1\}, X_i \in \{0, 1\}$ 

where  $M_i$  is randomly drawn from  $\{0, 1\}$  with equal probability for each individual *i*.  $y_i, X_i, D_i$  are observed variables. I experiment on the following two transformations of  $M_i$ :

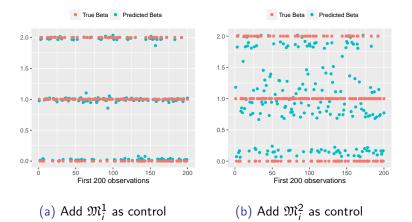
$$\mathfrak{M}_{i}^{1} = \left\{ \begin{array}{cc} -|\epsilon_{i}| & \text{if } M_{i} = 0 \\ |\epsilon_{i}| & \text{if } M_{i} = 1 \end{array} \right\}; \qquad \mathfrak{M}_{i}^{2} = \left\{ \begin{array}{cc} -|\epsilon_{i}| + X_{i} & \text{if } M_{i} = 0 \\ |\epsilon_{i}| + X_{i} & \text{if } M_{i} = 1 \end{array} \right\}$$

Where  $\epsilon_i \sim N(0, 1)$ 

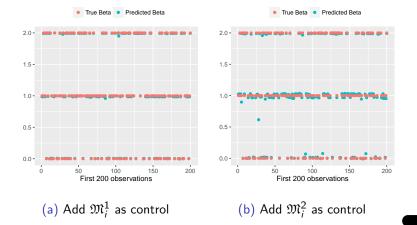
# Weak (a) and conditional weak (b) monotonic transformations of $M_i$



### Estimates, n = 200



### Estimates, n = 2000



### Two dimensional examples

$$\begin{split} p_{d,t} &= 10 + \beta_{d,t} e_{d,t} + ms_{d,t} + mc_t + \epsilon_{d,t} \\ ms_{d,t} &= u_{d,t} + 0.1 e_{d,t} \\ mc_t &= u_t - 0.1 \overline{e_t}; \quad \overline{e_t} := \frac{\sum_d e_{d,t}}{n_d} \\ e_{d,t} \sim N(0,1), \ u_{d,t} \sim \textit{Uniform}(0,1), \ \epsilon_{d,t} \sim N(0,0.01) \end{split}$$

Simulate 3 different cases of the underlying ERPT function

$$\begin{array}{ll} \textit{Spec1}: & \beta_{d,t} = (ms_{d,t} - 0.5)^2 + mc_t; \ u_t \sim \textit{Uniform}(0,1) \\ \textit{Spec2}: & \beta_{d,t} = 2(ms_{d,t} - 0.5)^2 * mc_t; \ u_t \sim \textit{Uniform}(0,1) \\ \textit{Spec3}: & \beta_{d,t} = 2(ms_{d,t} - 0.5)^2 * mc_t; \ u_t \sim \textit{N}(0,1) \\ \end{array}$$

Objective:

- $d, t, p_{d,t}, e_{d,t}, ms_{d,t}$  observed;  $mc_t$  unobserved.
- Estimate  $\beta_{d,t}$  and understand what explains the heterogeneity of  $\beta_{d,t}$ .

# It may be difficult to reach the correct regression specification, column (5)

Simulation Spec1: $p_{d,t} = \mu + \left[ (ms_{d,t} - 0.5)^2 + mc_t \right] e_{d,t} + ms_{d,t} + mc_t + \epsilon_{d,t}$					
	(1)	(2)	(3)	(4)	(5)
e <sub>d,t</sub>	0.757*** (0.002)	0.635*** (0.001)	0.759*** (0.003)	0.638*** (0.002)	0.796*** (0.003)
ms <sub>d,t</sub>	× ,	1.198*** (0.004)		1.198*** (0.004)	0.995*** (0.005)
$e_{d,t} * ms_{d,t}$		. ,	-0.006 (0.006)	-0.006 (0.004)	-1.019*** (0.011)
$e_{d,t} * ms_{d,t}^2$					1.016*** (0.011)
Time FE	yes	yes	yes	yes	yes
Individual FE Adjusted R <sup>2</sup> Observations	yes 0.663 80,000	yes 0.865 80,000	yes 0.722 80,000	yes 0.865 80,000	yes 0.879 80,000

Even at the correct regression specification column (5), results are not very informative about the underlying structure driving the heterogeneity of ERPT,  $\beta_{d,t}$ .

$$\begin{split} & \textit{Spec1:} \quad \beta_{d,t} = (ms_{d,t} - 0.5)^2 + mc_t; \ u_t \sim \textit{Uniform}(0,1) \\ & \textit{Spec2:} \quad \beta_{d,t} = 2(ms_{d,t} - 0.5)^2 * mc_t; \ u_t \sim \textit{Uniform}(0,1) \\ & \textit{Spec3:} \quad \beta_{d,t} = 2(ms_{d,t} - 0.5)^2 * mc_t; \ u_t \sim \textit{N}(0,1) \end{split}$$

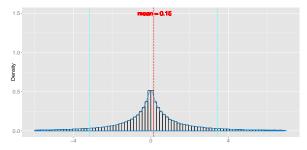
	Spec1	Spec2	Spec3
e <sub>d,t</sub>	0.796***	0.242***	-0.088***
	(0.003)	(0.001)	(0.019)
ms <sub>d.t</sub>	0.995***	1.002***	1.038***
	(0.005)	(0.002)	(0.016)
$e_{d,t} * ms_{d,t}$	-1.019* <sup>**</sup>	-0.962* <sup>**</sup>	0.301***
	(0.011)	(0.004)	(0.016)
$e_{d,t} * ms_{d,t}^2$	1.016***	0.957***	-0.190***
u,t	(0.011)	(0.004)	(0.011)
Time FE	yes	yes	yes
Individual FE	yes	yes	yes
Adjusted R <sup>2</sup>	0.879	0.885	0.056
Observations	80,000	80,000	80,000

Conventional estimation methods of ERPT trade-off between controlling unobserved variables and flexibility of functional forms

1 Interaction terms with e<sub>d,t</sub> and multiple fixed effects

 $p_{i,f,d,t} = \beta e_{d,t} + InteractionTerms + Controls + FEs + \epsilon_{i,f,d,t}$ 

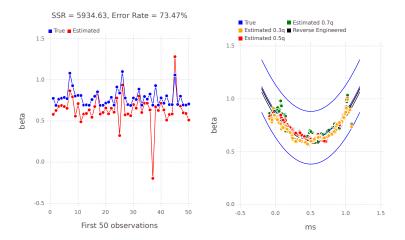
Or more flexible settings by categorizing data into several bins, e.g. by destinations, quantiles of market share, etc.



Distribution of Quarterly ERPT Estimates for China's Exporters by Firm-product Bins

### Misalignment of beta due to unobserved $mc_t$

$$p_{d,t} = 10 + \underbrace{\left[(ms_{d,t} - 0.5)^2 + mc_t\right]}_{beta} e_{d,t} + ms_{d,t} + mc_t + \epsilon_{d,t}$$



# **RHS** Graph

### True

- Plot  $f(ms) = (ms 0.5)^2 + mc$
- where  $mc^q := quantile(mc, q)$  from data;  $q \in [0.3, 0.5, 0.7]$
- 2 Estimated
  - Calculate q quantile of  $e_{d,t}$ ,  $b_t^0$ ,  $b_t^1$ ,  $b_d^0$ ,  $b_d^1$
  - Plot  $f(ms) = g_2((e_{d,t})^q, ms, (b_t^0)^q, (b_t^1)^q, (b_d^0)^q, (b_d^1)^q)$
- 8 Reverse Engineered

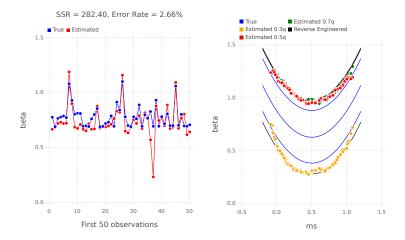
# **1** Use $mc^q := quantile(mc, q)$ from reverse engineering; $q \in [0.3, 0.5, 0.7]$

- Run GBRT with  $mc_t$  as the dependent variables on feature variables  $e_{d,t}$ ,  $b_t^0$ ,  $b_t^1$ ,  $b_d^0$ ,  $b_d^1$  and get  $M_3(.)$
- Estimate  $mc^q = M_3((e_{d,t})^q, (b_t^0)^q, (b_t^1)^q, (b_d^0)^q, (b_d^1)^q)$

2 Plot  $f(ms) = (ms - 0.5)^2 + mc$ 

# Adding Control Dummies

$$p_{d,t} = 10 + \underbrace{\left[(ms_{d,t} - 0.5)^2 + mc_t\right]}_{beta} e_{d,t} + ms_{d,t} + mc_t + \epsilon_{d,t}$$

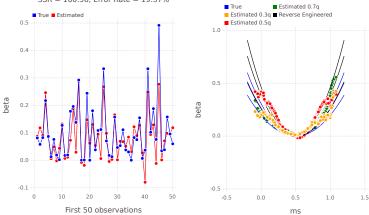


# Spec2: Misalignment of beta due to unobserved *mc*<sub>t</sub>

Estimated 0.7q True Estimated 0.3g Reverse Engineered True Estimated 0.6 Estimated 0.5g 0.7 0.6 0.4 0.5 0.2 0.4 beta oeta 0.3 0.0 0.2 0.1 -0.2 0.0 -0.4 -0.1 0 10 20 30 40 50 -0.5 0.0 0.5 1.0 First 50 observations ms

SSR = 817.50, Error Rate = 44.41%

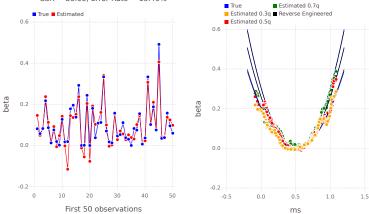
# Spec2: Adding Control Dummies



SSR = 160.56, Error Rate = 19.37%

Bac

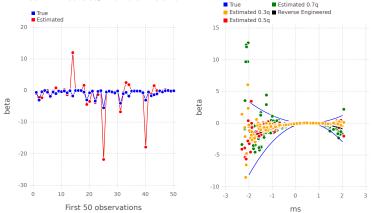
# Spec2: The Proposed Algorithm



SSR = 86.05, Error Rate = 13.40%

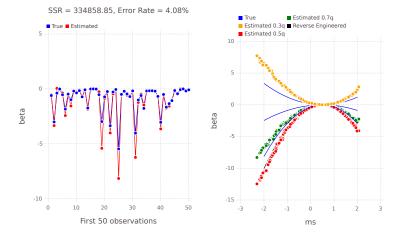
Back

# Spec3: Misalignment of beta due to unobserved *mc*<sub>t</sub>



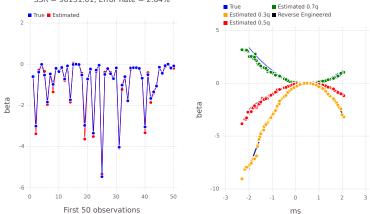
SSR = 1425815.68. Error Rate = 25.84%

# Spec3: Adding Control Dummies



Back

# Spec3: The Proposed Algorithm



SSR = 96151.01, Error Rate = 2.64%

Back

# High Nonlinearity

Setting:

$$p_{d,t} = 10 + \beta_{d,t}e_{d,t} + ms_{d,t} - mc_t + \epsilon_{d,t}$$
  

$$\beta_{d,t} = (ms_{d,t} - 0.5)^2 + sin(1000mc_t)mc_t$$
  

$$ms_{d,t} = u_{d,t} + 0.1e_{d,t}$$
  

$$mc_t = u_t - 0.1\overline{e_t}$$
  

$$\overline{e_t} = \frac{\sum_d e_{d,t}}{n_d}$$
  

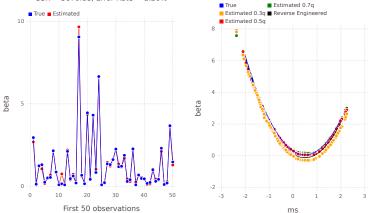
$$n_d = 2000; \ n_t = 40$$
  

$$u_{d,t} \sim N(0,1), \ u_t \sim N(0,1), \ \epsilon_{d,t} \sim N(0,1)$$

Objective:

- Read records d, t, p<sub>d,t</sub>, e<sub>d,t</sub>, ms<sub>d,t</sub>
- Estimate  $\beta_{d,t}$





SSR = 5670.88, Error Rate = 1.26%

#### Bacl

### Not Identifiable

### Setting:

$$p_{d,t} = 10 + \beta_{d,t}e_{d,t} + ms_{d,t} - mc_{d,t} + \epsilon_{d,t}$$
  

$$\beta_{d,t} = (ms_{d,t} - 0.5)^2 + mc_{d,t}$$
  

$$ms_{d,t} = u_{d,t} + 0.1e_{d,t}$$
  

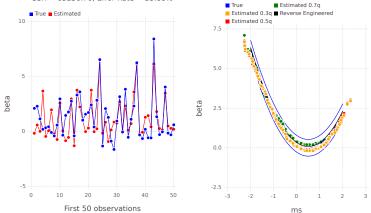
$$mc_{d,t} = u_{d,t} - 0.1e_{d,t}$$
  

$$n_d = 2000; \ n_t = 40$$
  

$$u_{d,t} \sim N(0,1), \ u_t \sim N(0,1), \ \epsilon_{d,t} \sim N(0,1)$$

- Read records d, t, p<sub>d,t</sub>, e<sub>d,t</sub>, ms<sub>d,t</sub>
- Estimate  $\beta_{d,t}$

### Not Identifiable My algorithm



SSR = 65236.70, Error Rate = 35.88%

## Larger Correlation

Setting:

$$p_{d,t} = 10 + \beta_{d,t}e_{d,t} + ms_{d,t} - mc_t + \epsilon_{d,t}$$
  

$$\beta_{d,t} = (ms_{d,t} - 0.5)^2 + mc_t$$
  

$$ms_{d,t} = u_{d,t} + 0.1e_{d,t}$$
  

$$mc_t = u_t - 1\overline{e_t}$$
  

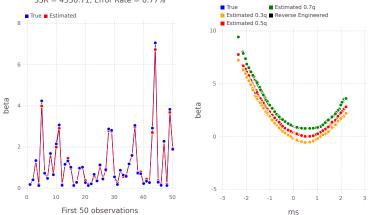
$$\overline{e_t} = \frac{\sum_d e_{d,t}}{n_d}$$
  

$$n_d = 2000; \ n_t = 40$$
  

$$u_{d,t} \sim N(0,1), \ u_t \sim N(0,1), \ \epsilon_{d,t} \sim N(0,1)$$

- Read records d, t, p<sub>d,t</sub>, e<sub>d,t</sub>, ms<sub>d,t</sub>
- Estimate  $\beta_{d,t}$





SSR = 4330.71, Error Rate = 0.77%

## Different Function of the Outer Part

Setting:

$$p_{d,t} = 10 + \beta_{d,t}e_{d,t} + ms_{d,t} - mc_t^2 + \epsilon_{d,t}$$
  

$$\beta_{d,t} = (ms_{d,t} - 0.5)^2 + mc_t$$
  

$$ms_{d,t} = u_{d,t} + 0.1e_{d,t}$$
  

$$mc_t = u_t - 0.1\overline{e_t}$$
  

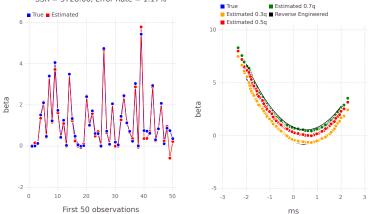
$$\overline{e_t} = \frac{\sum_d e_{d,t}}{n_d}$$
  

$$n_d = 2000; \ n_t = 40$$
  

$$u_{d,t} \sim N(0,1), \ u_t \sim N(0,1), \ \epsilon_{d,t} \sim N(0,1)$$

- Read records d, t, p<sub>d,t</sub>, e<sub>d,t</sub>, ms<sub>d,t</sub>
- Estimate  $\beta_{d,t}$

### Different Function of the Outer Part My algorithm



SSR = 5728.66, Error Rate = 1.17%

### Arellano and Bond

Setting:

$$p_{d,t} = 0.95 p_{d,t-1} + \beta_{d,t} e_{d,t} + ms_{d,t} - mc_t$$
  

$$\beta_{d,t} = (ms_{d,t} - 0.5)^2 + mc_t$$
  

$$ms_{d,t} = u_{d,t} + 0.1 e_{d,t}$$
  

$$mc_t = u_t - 0.1\overline{e_t}$$
  

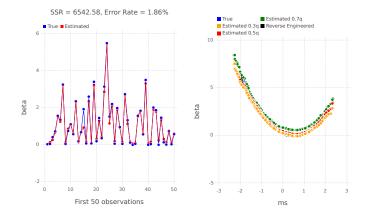
$$\overline{e_t} = \frac{\sum_d e_{d,t}}{n_d}$$
  

$$n_d = 2000; \ n_t = 40$$
  

$$u_{d,t} \sim N(0,1), \ u_t \sim N(0,1), \ \epsilon_{d,t} \sim N(0,1)$$

- Read records d, t, p<sub>d,t</sub>, e<sub>d,t</sub>, ms<sub>d,t</sub>
- Estimate  $\beta_{d,t}$

### Arellano and Bond My Algorithm



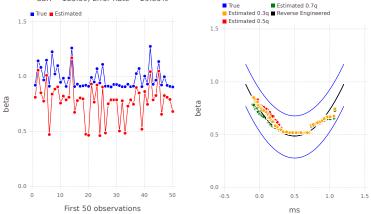
### Reduce Sample Size

Setting:

$$\begin{aligned} p_{d,t} &= 10 + \beta_{d,t} e_{d,t} + ms_{d,t} - mc_t + \epsilon_{d,t} \\ \beta_{d,t} &= (ms_{d,t} - 0.5)^2 + mc_t \\ ms_{d,t} &= u_{d,t} + 0.1 e_{d,t} \\ mc_t &= u_t - 0.1 \overline{e_t} \\ \overline{e_t} &= \frac{\sum_d e_{d,t}}{n_d} \\ n_d &= 200; \ n_t = 40 \\ u_{d,t} \sim \textit{uniform, } u_t \sim \textit{uniform, } \epsilon_{d,t} \sim \textit{N}(0,1) \end{aligned}$$

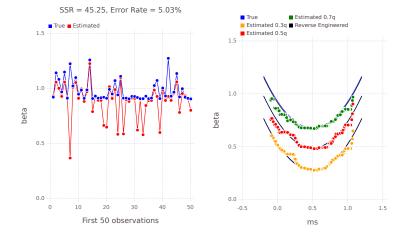
- Read records d, t, p<sub>d,t</sub>, e<sub>d,t</sub>, ms<sub>d,t</sub>
- Estimate  $\beta_{d,t}$

# Reduce Sample Size

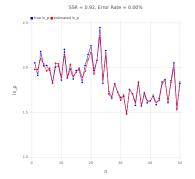


SSR = 183.59, Error Rate = 29.35%

### Reduce Sample Size My Algorithm



## Precision on Estimating Price



SSR = 0.03. Error Rate = 0.00%

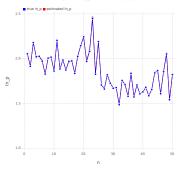
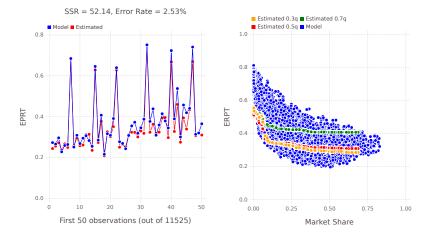


Figure 6: Decision Tree GBM

Figure 7: My Algorithm

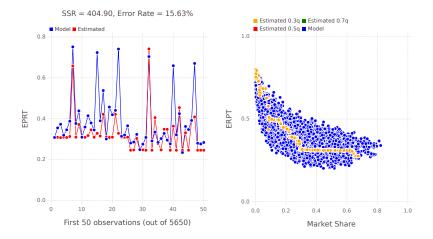
# Case 1: Only Exchange Rate Shocks

My algorithm compared to true counter-factual environments



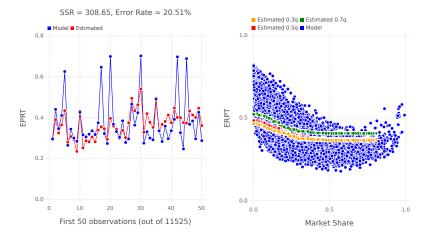
# Case 1: Only Exchange Rate Shocks

Without adding regression coefficients



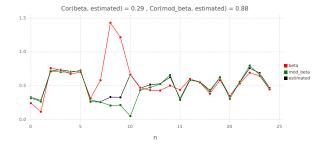
# Case 2: Add Productivity Shocks

My algorithm compared to true counter-factual environments



## Recover ERPT for Simulated Exporters

Case 3: add productivity shocks to the exporting country



Firm's productivity is assumed to follow a AR(1) process with a persistence of  $\rho = 0.95$ . The red line presents the ERPT estimates calculated using actual price changes of the simulated model. The green line represents the model implied ERPT estimates in a counter-factual equilibrium where there is no productivity shock in the next period. The black line represents ERPT estimates predicted by the algorithm.

# More Experiments

High nonlinear function of the unobserved variable



- 2 Large correlation between the unobserved variable and exchange rates 🖸
- 3 Not identifiable case , unobserved variables vary over all dimensions 💽
- 4 Arellano and Bond dynamic panel
- 6 Reduce sample size

# Algorithm

Input data I, y, X, e

- 1: Obtain variable names of the index matrix I and the feature variable matrix X and save them as  $i_{names}$  and  $x_{names}$  respectively.
- 2: Calculate all non-repetitive combinations of dimension indices in  $i_{names}$  and save as  $S_i$ .
- 3: for s in S; do  $\mathbf{I}_{s} \leftarrow \mathbf{I}[i_{names} \in s]$ 4:  $\widetilde{\mathbf{I}}_{s} \leftarrow unique(\mathbf{I}_{e})$ 5: 6: for x in x<sub>names</sub> do 7:  $x_{s} \leftarrow \mathbf{0}$ 8: for  $i_s$  in 1 to nrow( $\tilde{I}_s$ ) do  $x_{s}[\mathbf{I}_{s} = \widetilde{\mathbf{I}}[i_{s}]] \leftarrow mean(x|\mathbf{I}_{s} = \widetilde{\mathbf{I}}[i_{s}])$ 9: 10: end for 11: end for 12: end for 13: Calculate all non-repetitive binary combinations of  $S_i$  and save as  $S_{share}$ . 14: for s in S<sub>share</sub> do  $(s_a, s_b) \leftarrow s[sort(length(s[1], s[2]))]$ 15: 16: for x in x<sub>names</sub> do 17:  $x_{s_a,s_b} \leftarrow \frac{x_{s_a}}{x_{s_a}}$
- 18: end for
- 19: end for

# Algorithm (cont.)

- 20: Observe dimensions the supervisor y and the policy/treatment variable e vary. Identify a subset available for controlling unobserved variables and save as  $S_{id}$ . 21: for s in S<sub>id</sub> do 22: Assume a possible (linear) structural equation based on economic rationale. 23: for *j* in 1:(number of parameters in the structural model) do  $coef_{s}^{j} \leftarrow \mathbf{0}$ 24: 25: end for for  $d_s$  in 1 to nrow( $\tilde{\mathbf{I}}_s$ ) do 26: 27: Estimate the structural regression for the subset of data where  $I_s = \widetilde{I}[i_s]$ 28: for *j* in 1:(number of parameters in the structural model) do  $coef_{s}^{j}[\mathbf{I}_{s} = \widetilde{\mathbf{I}}[i_{s}]] \leftarrow parameter^{j}$ 29: 30: end for
- 31: end for
- 32: end for
- 33: Run GBRT with supervisor **y** on **e**, **X**,  $\mathbf{X}_{\mathbf{s}_{a},\mathbf{s}_{b}}$ , **coef**<sup>j</sup><sub>id</sub> and obtain model  $g_{1}$ .

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34: y^{\text{Est1}} \leftarrow g_1(e, X, X_s, \text{coef}_s^j)

35: y^{\text{Est2}} \leftarrow g_1(e + std(e), X, X_s, \text{coef}_s^j)

36: beta^{\text{Est}} \leftarrow \frac{y^{\text{Est2}} - y^{\text{Est1}}}{std(e)}

37: Run GBRT again with supervisor beta^{\text{Est}} on e, X, X_{s_a, s_b}, \text{coef}_{id}^j and obtain model g_2.

Output: g_1, g_2, beta^{\text{Est}} \bullet \text{Back}
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